Hidden Markov Models

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Overview

- Markov models
- Hidden Markov models

Markov Sequences

- ▶ Consider a sequence of random variables *X*₁*, X*₂*,...,X*_m where *m* is the length of the sequence
- Each variable X_i can take any value in $\{1, 2, \ldots, k\}$
- How do we model the joint distribution

$$P(X_1 = x_1, X_2 = x_2, \dots, X_m = x_m)$$

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The Markov Assumption

$$P(X_{1} = x_{1}, X_{2} = x_{2}, \dots, X_{m} = x_{m})$$

$$= P(X_{1} = x_{1}) \prod_{j=2}^{m} P(X_{j} = x_{j} | X_{1} = x_{1}, \dots, X_{j-1} = x_{j-1})$$

$$= P(X_{1} = x_{1}) \prod_{j=2}^{m} P(X_{j} = x_{j} | X_{j-1} = x_{j-1})$$

- The first equality is exact (by the chain rule).
- The second equality follows from the Markov assumption: for all j = 2...m,

$$P(X_j = x_j | X_1 = x_1, \dots, X_{j-1} = x_{j-1}) = P(X_j = x_j | X_{j-1} = x_{j-1})$$

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Homogeneous Markov Chains

In a homogeneous Markov chain, we make an additional assumption, that for j = 2...m,

$$P(X_j = x_j | X_{j-1} = x_{j-1}) = q(x_j | x_{j-1})$$

where q(x'|x) is some function

 Idea behind this assumption: the transition probabilities do not depend on the position in the Markov chain (do not depend on the index j)

Markov Models

Our model is then as follows:

$$p(x_1, x_2, \dots, x_m; \underline{\theta}) = q(x_1) \prod_{j=2}^m q(x_j | x_{j-1})$$

► Parameters in the model:

•
$$q(x)$$
 for $x = \{1, 2, \dots, k\}$
Constraints: $q(x) \ge 0$ and $\sum_{x=1}^{k} q(x) = 1$

•
$$q(x'|x)$$
 for $x = \{1, 2, ..., k\}$ and $x' = \{1, 2, ..., k\}$
Constraints: $q(x'|x) \ge 0$ and $\sum_{x'=1}^{k} q(x'|x) = 1$

A Generative Story for Markov Models

- ► A sequence x₁, x₂, ..., x_m is generated by the following process:
 - 1. Pick x_1 at random from the distribution q(x)

2. For
$$j = 2 \dots m$$
:

• Choose x_j at random from the distribution $q(x|x_{j-1})$

Today's Lecture

- Markov models
- Hidden Markov models

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Modeling Pairs of Sequences

- In many applications, we need to model pairs of sequences
- Examples:
 - 1. Part-of-speech tagging in natural language processing (assign each word in a sentence to one of the categories noun, verb, preposition etc.)
 - 2. Speech recognition (map acoustic sequences to sequences of words)
 - 3. Computational biology: recover gene boundaries in DNA sequences

Probabilistic Models for Sequence Pairs

- ▶ We have two sequences of random variables: X₁, X₂,..., X_m and S₁, S₂,..., S_m
- Intuitively, each X_i corresponds to an "observation" and each S_i corresponds to an underlying "state" that generated the observation. Assume that each S_i is in {1, 2, ... k}, and each X_i is in {1, 2, ... o}
- How do we model the joint distribution

$$P(X_1 = x_1, \dots, X_m = x_m, S_1 = s_1, \dots, S_m = s_m)$$

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Hidden Markov Models (HMMs)

► In HMMs, we assume that:

$$P(X_1 = x_1, \dots, X_m = x_m, S_1 = s_1, \dots, S_m = s_m)$$

= $P(S_1 = s_1) \prod_{j=2}^m P(S_j = s_j | S_{j-1} = s_{j-1}) \prod_{j=1}^m P(X_j = x_j | S_j = s_j)$

Independence Assumptions in HMMs

By the chain rule, the following equality is exact:

$$P(X_1 = x_1, \dots, X_m = x_m, S_1 = s_1, \dots, S_m = s_m)$$

$$= P(S_1 = s_1, \dots, S_m = s_m) \times P(X_1 = x_1, \dots, X_m = x_m | S_1 = s_1, \dots, S_m = s_m)$$

Assumption 1: the state sequence forms a Markov chain

$$P(S_1 = s_1, \dots, S_m = s_m) = P(S_1 = s_1) \prod_{j=2}^m P(S_j = s_j | S_{j-1} = s_{j-1})$$

Independence Assumptions in HMMs

By the chain rule, the following equality is exact:

$$P(X_1 = x_1, \dots, X_m = x_m | S_1 = s_1, \dots, S_m = s_m)$$

= $\prod_{j=1}^m P(X_j = x_j | S_1 = s_1, \dots, S_m = s_m, X_1 = x_1, \dots, X_{j-1} = x_j)$

 Assumption 2: each observation depends only on the underlying state

$$P(X_j = x_j | S_1 = s_1, \dots, S_m = s_m, X_1 = x_1, \dots, X_{j-1} = x_j)$$

= $P(X_j = x_j | S_j = s_j)$

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The Model Form for HMMs

The model takes the following form:

$$p(x_1 \dots x_m, s_1 \dots s_m; \underline{\theta}) = t(s_1) \prod_{j=2}^m t(s_j | s_{j-1}) \prod_{j=1}^m e(x_j | s_j)$$

- Parameters in the model:
 - 1. Initial state parameters t(s) for $s \in \{1, 2, \dots, k\}$
 - 2. Transition parameters t(s'|s) for $s, s' \in \{1, 2, \dots, k\}$
 - 3. Emission parameters e(x|s) for $s \in \{1, 2, \dots, k\}$ and $x \in \{1, 2, \dots, o\}$

A Generative Story for Hidden Markov Models

- ▶ Sequence pairs s₁, s₂,..., s_m and x₁, x₂,..., x_m are generated by the following process:
 - 1. Pick s_1 at random from the distribution t(s). Pick x_1 from the distribution $e(x|s_1)$

2. For
$$j = 2 \dots m$$
:

- Choose s_j at random from the distribution $t(s|s_{j-1})$
- Choose x_j at random from the distribution $e(x|s_j)$

Today's Lecture

- More on Hidden Markov models:
 - parameter estimation
 - The Viterbi algorithm

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Parameter Estimation with Fully Observed Data

▶ We'll now discuss parameter estimates in the case of fully observed data: for i = 1...n, we have pairs of sequences x_{i,j} for j = 1...m and s_{i,j} for j = 1...m. (i.e., we have n training examples, each of length m.)

Parameter Estimation: Transition Parameters

- ► Assume we have fully observed data: for i = 1...n, we have pairs of sequences x_{i,j} for j = 1...m and s_{i,j} for j = 1...m
- ▶ Define count(i, s → s') to be the number of times state s' follows state s in the i'th training example. More formally:

$$\mathsf{count}(i, s \to s') = \sum_{j=1}^{m-1} [[s_{i,j} = s \land s_{i,j+1} = s']]$$

(We define $[[\pi]]$ to be 1 if π is true, 0 otherwise.)

The maximum-likelihood estimates of transition probabilities are then

$$t(s'|s) = \frac{\sum_{i=1}^{n} \operatorname{count}(i, s \to s')}{\sum_{i=1}^{n} \sum_{s'} \operatorname{count}(i, s \to s')}$$

Parameter Estimation: Emission Parameters

- ► Assume we have fully observed data: for i = 1...n, we have pairs of sequences x_{i,j} for j = 1...m and s_{i,j} for j = 1...m
- ▶ Define count(i, s → x) to be the number of times state s is paired with emission x. More formally:

$$\mathsf{count}(i, s \rightsquigarrow x) = \sum_{j=1}^{m} [[s_{i,j} = s \land x_{i,j} = x]]$$

The maximum-likelihood estimates of emission probabilities are then

$$e(x|s) = \frac{\sum_{i=1}^{n} \operatorname{count}(i, s \rightsquigarrow x)}{\sum_{i=1}^{n} \sum_{x} \operatorname{count}(i, s \rightsquigarrow x)}$$

Parameter Estimation: Initial State Parameters

- ► Assume we have fully observed data: for i = 1...n, we have pairs of sequences x_{i,j} for j = 1...m and s_{i,j} for j = 1...m
- Define count(i, s) to be 1 if state s is the initial state in the sequence, and 0 otherwise:

$$\mathsf{count}(i,s) = [[s_{i,1} = s]]$$

The maximum-likelihood estimates of initial state probabilities are:

$$t(s) = \frac{\sum_{i=1}^{n} \operatorname{count}(i, s)}{n}$$

Today's Lecture

- Hidden Markov models:
 - parameter estimation
 - the Viterbi algorithm

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The Viterbi Algorithm

• Goal: for a given input sequence x_1, \ldots, x_m , find

$$\arg\max_{s_1,\ldots,s_m} p(x_1\ldots x_m, s_1\ldots s_m; \underline{\theta})$$

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► This is the most likely state sequence s₁...s_m for the given input sequence x₁...x_m

The Viterbi Algorithm

• Goal: for a given input sequence x_1, \ldots, x_m , find

$$\arg\max_{s_1,\ldots,s_m} p(x_1\ldots x_m, s_1\ldots s_m; \underline{\theta})$$

The Viterbi algorithm is a dynamic programming algorithm. Basic data structure:

$$\pi[j,s]$$

will be a table entry that stores the maximum probability for any state sequence ending in state s at position j. More formally: $\pi[1, s] = t(s)e(x_1|s)$, and for j > 1,

$$\pi[j,s] = \max_{s_1\dots s_{j-1}} \left[t(s_1)e(x_1|s_1) \left(\prod_{k=2}^{j-1} t(s_k|s_{k-1})e(x_k|s_k) \right) t(s|s_{j-1})e(x_j|s) \right]$$

The Viterbi Algorithm

• Initialization: for
$$s = 1 \dots k$$

$$\pi[1,s] = t(s)e(x_1|s)$$

► For
$$j = 2...m$$
, $s = 1...k$:

$$\pi[j, s] = \max_{s' \in \{1...k\}} [\pi[j - 1, s'] \times t(s|s') \times e(x_j|s)]$$

We then have

$$\max_{s_1\dots s_m} p(x_1\dots x_m, s_1\dots s_m; \underline{\theta}) = \max_s \pi[m, s]$$

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• The algorithm runs in $O(mk^2)$ time

Viterbi as a Shortest-Path Algorithm

- The input sequence $x_1 \dots x_m$ is fixed
- ▶ Have vertices in a graph labeled (j, s) for $s \in \{1 ... k\}$ and j = 1 ... m. In addition have a source vertex labeled 0
- For s ∈ {1...k}, we have a directed edge from vertex 0 to vertex (1,s), with weight t(s)e(x₁|s)
- ▶ For each j = 2...m, and $s, s' \in \{1...k\}$, have a directed edge from (j 1, s) to (j, s') with weight $t(s'|s)e(x_j|s')$ (the weight of any path is the product of weights on edges in the path)
- ► π[j,s] is the highest weight for any path from vertex 0 to vertex (j,s)

The Viterbi Algorithm: Backpointers

• Initialization: for
$$s = 1 \dots k$$

$$\pi[1,s] = t(s)e(x_1|s)$$

► For
$$j = 2...m$$
, $s = 1...k$:

$$\pi[j, s] = \max_{s' \in \{1...k\}} [\pi[j - 1, s'] \times t(s|s') \times e(x_j|s)]$$
and

$$bp[j,s] = \arg\max_{s' \in \{1...k\}} [\pi[j-1,s'] \times t(s|s') \times e(x_j|s)]$$

The bp entries are backpointers that will allow us to recover the identity of the highest probability state sequence Viterbi Algorithm: Backpointers (continued)

Highest probability for any sequence of states is

 $\max_s \pi[m,s]$

To recover identity of highest-probability sequence:

$$s_m = \arg\max_s \pi[m, s]$$

and for $j = m \dots 2$,

$$s_{j-1} = bp[j, s_j]$$

• The sequence of states $s_1 \dots s_m$ is then

$$\arg\max_{s_1,\ldots,s_m} p(x_1\ldots x_m, s_1\ldots s_m; \underline{\theta})$$