

# Linear regression without correspondence

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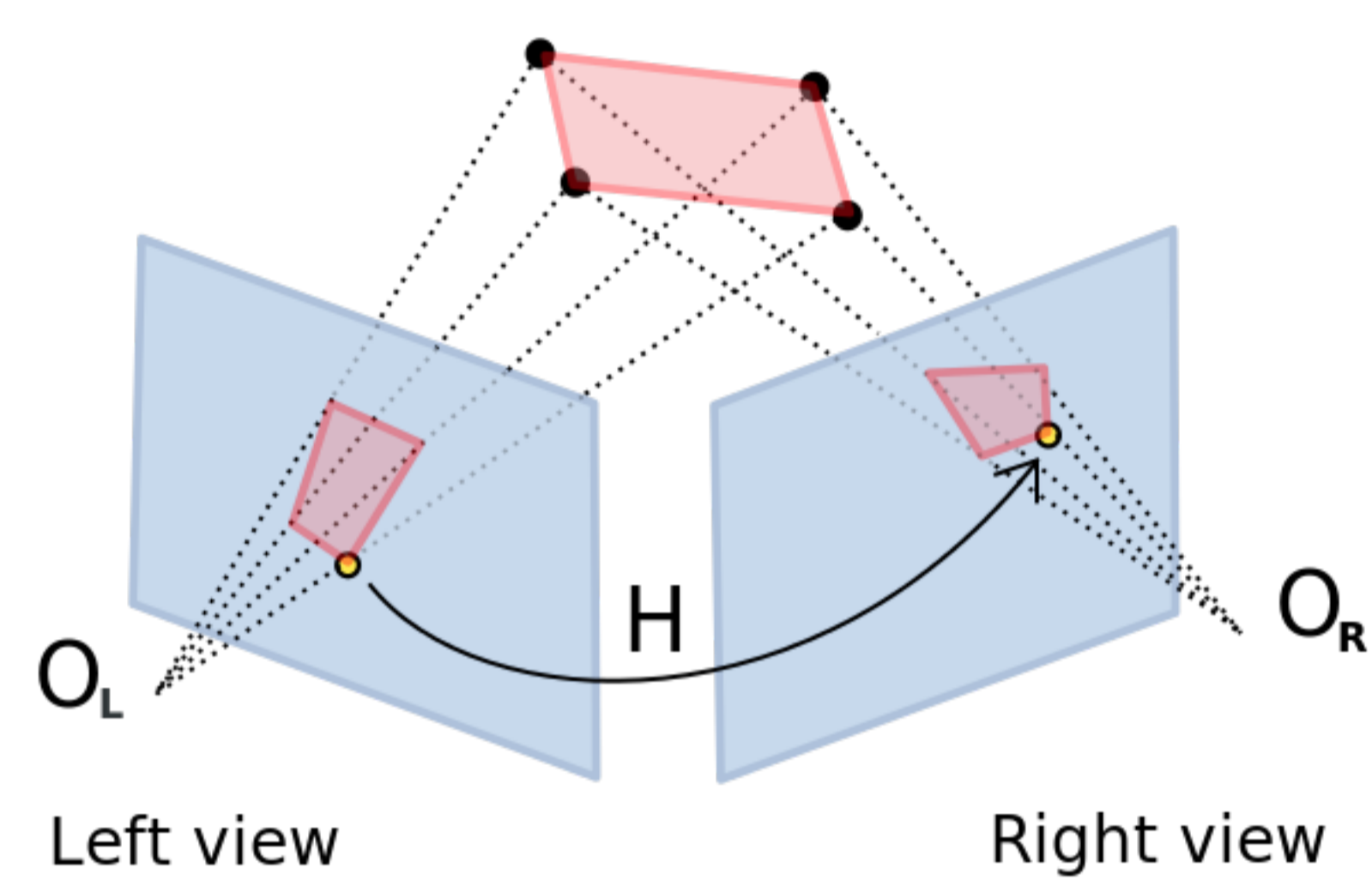
## Problem definition

- ▷ **Covariate vectors:**  $x_1, x_2, \dots, x_n \in \mathbb{R}^d$
- ▷ **Responses:**  $y_1, y_2, \dots, y_n \in \mathbb{R}$
- ▷ **Model:**

$$y_i = \bar{w}^\top x_{\bar{\pi}(i)} + \varepsilon_i, \quad i \in [n]$$
- ▷ Unknown linear function:  $\bar{w} \in \mathbb{R}^d$
- ▷ Unknown permutation:  $\bar{\pi} \in S_n$
- ▷ Measurement errors:  $\varepsilon_1, \varepsilon_2, \dots, \varepsilon_n \in \mathbb{R}$   
e.g.,  $(\varepsilon_i)_{i=1}^n$  iid from  $\mathcal{N}(0, \sigma^2)$

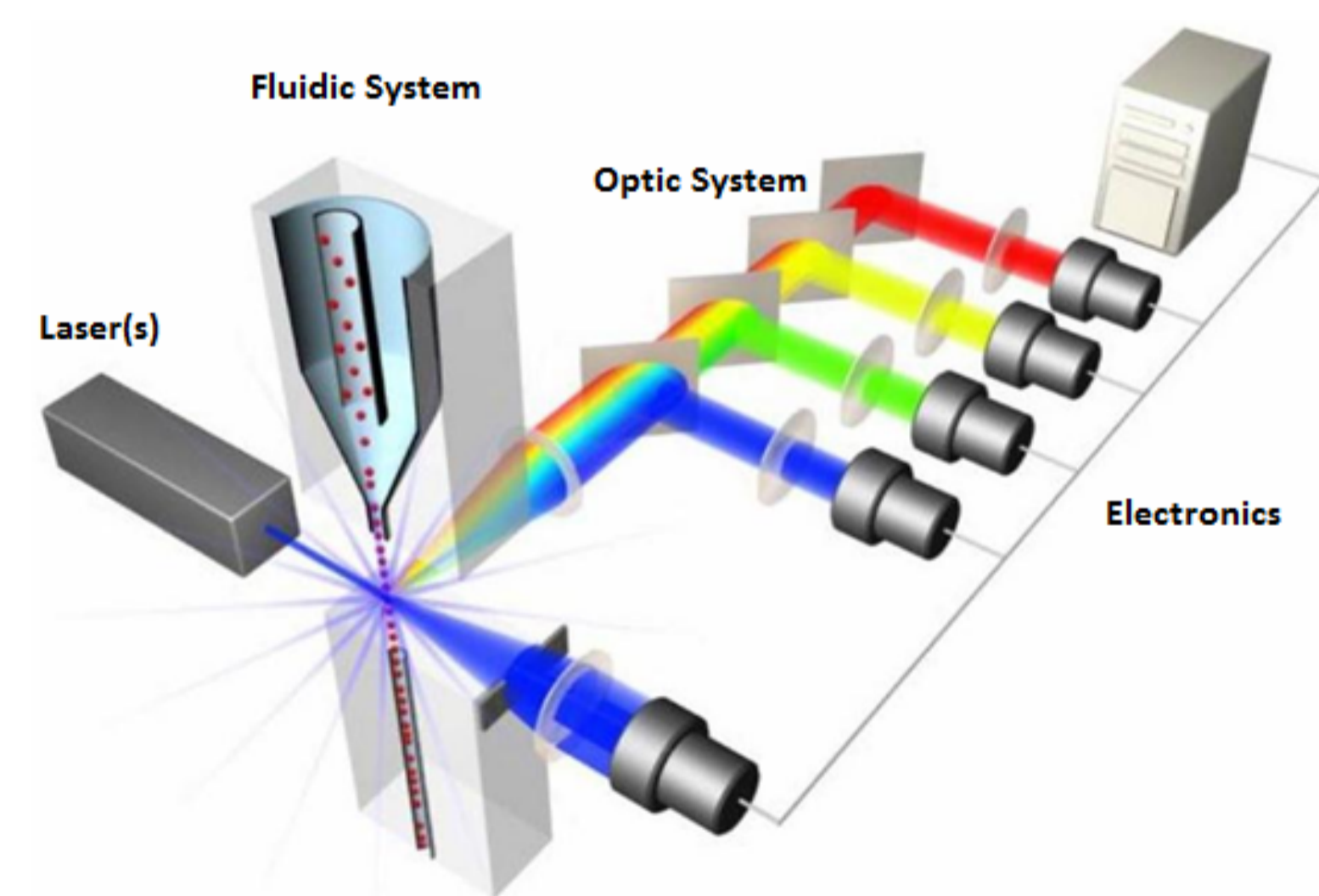
## Examples

### Multi-view geometry



- ▷ Unknown correspondence between keypoints

### Flow cytometry



- ▷ Observe the entire emission spectrum at once

## Strong NP-hardness

### Definition 1 (Permuted Linear System).

Given  $X \in \mathbb{Z}^{n \times d}$ ,  $Y \in \mathbb{Q}^n$ , decide if there exists a vector  $w \in \mathbb{Q}^d$  and a permutation  $\pi \in S_n$  such that  $Xw = Y_\pi$

**Proposition 1.** Permuted Linear System is strongly NP-complete by a reduction from 3-Partition.

## Approximation guarantee for least-squares

### Definition 2 (Least-squares recovery).

Given  $(x_i)_{i=1}^n$  and  $(y_i)_{i=1}^n$  from  $\mathbb{R}$ , find

$$(\hat{w}_{mle}, \hat{\pi}_{mle}) := \arg \min_{w \in \mathbb{R}, \pi \in S_n} \sum_{i=1}^n (y_i - wx_{\pi(i)})^2$$

**Theorem 1.** There is an algorithm that given any inputs  $(x_i)_{i=1}^n$ ,  $(y_i)_{i=1}^n$ , and  $\epsilon \in (0, 1)$ , returns a  $(1 + \epsilon)$ -approximate solution to the least squares problem in time  $(n/\epsilon)^{O(k)} + \text{poly}(n, d)$ , where  $k = \dim(\text{span}(x_i)_{i=1}^n)$ .

## Approximation algorithm

This uses the following coresnet result for linear systems:

### Proposition 2 (Boutsidis, Drineas, Magdon-Ismail).

Given a matrix  $A \in \mathbb{R}^{n \times k}$ , there exists a weighted subset of  $4k$  rows determined by a matrix  $S \in \mathbb{R}^{4k \times n}$  such that for any  $b$ , every minimizer of the subsampled linear system

$$w' \in \arg \min_w \|S(Aw - b)\|_2^2$$

also satisfies

$$\|Aw' - b\|_2^2 \leq c$$

for  $c = O(n/k)$ . Moreover, there exists an efficient algorithm which returns a matrix  $S$  in time  $\text{poly}(n, k)$ .

### Algorithm 1 Approximation algorithm

**input** Covariate matrix  $X = [x_1 | x_2 | \dots | x_n]^\top \in \mathbb{R}^{n \times k}$ ; response vector  $y = (y_1, y_2, \dots, y_n)^\top \in \mathbb{R}^n$ ; approximation parameter  $\epsilon \in (0, 1)$ .

- 1: Compute the matrix  $S \in \mathbb{R}^{4k \times n}$  from input matrix  $X$ .
- 2: Let  $\mathcal{B}$  be the set of all permutations of  $y$
- 3: Let  $c := 1 + 4(1 + \sqrt{n/(4k)})^2$ .
- 4: **for each**  $b \in \mathcal{B}$  **do**
- 5: Compute  $\tilde{w}_b \in \arg \min_{w \in \mathbb{R}^k} \| [0] S(Xw - b) \|_2^2$ , and let  $r_b := \min_{\Pi \in \mathcal{P}_n} \| [0] X \tilde{w}_b - \Pi^\top y \|_2^2$ .
- 6: Construct a  $\sqrt{\epsilon r_b}/c$ -net  $\mathcal{N}_b$  for the Euclidean ball of radius  $\sqrt{\epsilon r_b}$  around  $\tilde{w}_b$ , so that for each  $v \in \mathbb{R}^k$  with  $\| [0] v - \tilde{w}_b \|_2 \leq \sqrt{\epsilon r_b}$ , there exists  $v' \in \mathcal{N}_b$  such that  $\| [0] v - v' \|_2 \leq \sqrt{\epsilon r_b}/c$ .
- 7: **end for**
- 8: **return**

$$\hat{w} \in \arg \min_{w \in \cup_{b \in \mathcal{B}} \mathcal{N}_b} \min_{\Pi \in \mathcal{P}_n} \|Xw - \Pi^\top y\|_2^2$$

and

$$\hat{\Pi} \in \arg \min_{\Pi \in \mathcal{P}_n} \|X\hat{w} - \Pi^\top y\|_2^2$$

## Polynomial time recovery in the random setting

**Theorem 2.** Fix any  $\bar{w} \in \mathbb{R}^d$  and  $\bar{\pi} \in S_n$ , and assume  $n \geq d$ . Suppose  $(x_i)_{i=0}^n$  are drawn iid from  $\mathcal{N}(0, I_d)$ , and  $(y_i)_{i=0}^n$  satisfy

$$y_0 = \bar{w}^\top x_0; \quad y_i = \bar{w}^\top x_{\bar{\pi}(i)}, \quad i \in [n].$$

There is an algorithm that, given inputs  $(x_i)_{i=0}^n$  and  $(y_i)_{i=0}^n$ , returns  $\bar{\pi}$  and  $\bar{w}$  with high probability.

## Reduction to (random) subset sum

Given  $d + 1$  measurements and one correspondence  $y_0 = \bar{w}^\top x_0$ , for orthogonal  $(x_i)_{i=0}^n$ , can write:

$$\begin{aligned} y_0 &= \sum_{j=1}^d (\bar{w}^\top x_j) (x_j^\top x_0) = \sum_{j=1}^d y_{\bar{\pi}^{-1}(j)} (x_j^\top x_0) \\ &= \sum_{i=1}^d \sum_{j=1}^d \mathbb{1}\{\bar{\pi}(i) = j\} \cdot \underbrace{y_i (x_j^\top x_0)}_{c_{i,j}} \end{aligned}$$

- ▷  $\{c_{i,j}\}$  and  $y_0$  define a subset sum problem whose solution recovers the underlying correspondence.
- ▷ In general  $(x_i)_{i=0}^n$  are close to orthogonal; use the Moore-Penrose pseudoinverse.
- ▷ The one given correspondence can be brute-forced, creating  $d + 1$  subset sum instances of which only one has a solution

## Solving random subset-sum instances

### Proposition 3 (Lagarias and Odlyzko).

Random instances of subset sum are efficiently solvable when the  $c_{i,j}$ 's are independently and uniformly distributed over a large enough subinterval of  $\mathbb{Z}$ .

This relies on the following inequality which lower bounds the closeness to the target sum of incorrect solutions.

**Lemma 1.** For any vector  $(z_{i,j})$  which is not the correct correspondence,

$$\left| y_0 - \sum_{i,j} z_{i,j} c_{i,j} \right| \geq \frac{1}{2^{\text{poly}(d)}} \|\bar{w}\|_2$$

- ▷ We show this bound holds under other distributions satisfying general anticoncentration bounds and even if the  $c_{i,j}$ 's are not independent

## Reduction to shortest vector problem

### Definition 3 (Shortest vector problem).

Given a lattice basis  $B \subset \mathbb{R}^d$ , output a lattice vector  $Bz \in \Lambda B$  where

$$z = \arg \min_{z \in \mathbb{Z} - \{0\}} \|Bz\|_2^2$$

### Lemma 2 (LLL Lattice Basis Reduction).

There is an efficient approximation algorithm for solving the Shortest Vector Problem with

- ▷ Approximation factor:  $2^{d/2}$
- ▷ Running time:  $\text{poly}(d, \log \lambda(B))$

### Algorithm 2 Lattice algorithm for subset sum

**input** Source numbers  $\{c_i\}_{i \in \mathcal{I}} \subset \mathbb{R}$ ; target sum  $t \in \mathbb{R}$ ; lattice parameter  $\beta > 0$ .

- 1: Construct lattice basis  $B \in \mathbb{R}^{(|\mathcal{I}|+2) \times (|\mathcal{I}|+1)}$  where

$$B := \begin{bmatrix} I_{|\mathcal{I}|+1} \\ \beta t | -\beta c_i : i \in \mathcal{I} \end{bmatrix} \in \mathbb{R}^{(|\mathcal{I}|+2) \times (|\mathcal{I}|+1)}.$$

- 2: Run LLL Lattice Basis Reduction to find non-zero lattice vector  $v$  of length at most  $2^{|\mathcal{I}|/2} \cdot \lambda_1(B)$ .

## Information-theoretic lower bounds on SNR

### Definition 4 (Random measurement setting).

Observe

$$y_i = \bar{w}^\top x_{\bar{\pi}(i)} + \epsilon_i$$

where

- ▷  $\epsilon_i \stackrel{iid}{\sim} \mathcal{N}(0, \sigma^2)$  is the measurement noise
- ▷  $x_i \stackrel{iid}{\sim} \mathcal{N}(0, I_d)$  are the covariates

### Definition 5 (SNR).

The signal-to-noise ratio for this model is  $\|\bar{w}\|_2^2 / \sigma^2$

**Theorem 3.** In the random measurement setting, if for some constant  $C$

$$\text{SNR} \leq C \cdot \min \left\{ \frac{d}{\log \log n}, 1 \right\}$$

then for every estimator  $\hat{w}$ , there exists a  $\bar{w} \in \mathbb{R}^d$  such that

$$\mathbb{E} [\|\hat{w} - \bar{w}\|_2] \geq \frac{1}{24} \|\bar{w}\|_2$$

- ▷ Recall that standard linear regression satisfies the bound  $\mathbb{E} [\|w_{mle} - \bar{w}\|_2] \leq C\sigma\sqrt{d/n}$
- ▷ In the low SNR regime, more measurements makes the problem more difficult