

Majorization for CRFs and Latent Likelihoods

Tony Jebara and Anna Choromanska Columbia University

February 18, 2013

◆□▶ ◆□▶ ◆三▶ ◆三▶ 三回 ● のへで

Machine Learning

Unsupervised

- Given $\mathcal{D} = (y_t)_{t=1}^T$ sampled *iid* from unknown P(y)
- Given family of functions $p_{\theta}(y)$ parametrized by θ
- Find θ by minimizing some cost (e.g. partition function)
- Output $p_{\theta}(y)$
- Supervised
 - Given $\mathcal{D} = (x_t, y_t)_{t=1}^T$ sampled *iid* from unknown P(x, y)
 - Given family of functions $p_{\theta}(y|x)$ parametrized by θ
 - Find θ by minimizing some cost (e.g. partition function)
 - Predict using $\hat{y} = \arg \max_{y} p_{\theta}(y|x)$
- Today: a new simple bound for such optimizations

Optimization in Learning: Three Schools of Thought

Ways to optimize parameters to data

- First order methods
 - Steepest descent
 - Conjugate gradient
 - Stochastic gradient descent
 - Bundle methods
- Second order methods
 - Newton
 - BFGS [Broyden; Fletcher; Goldfarb; Shanno '70]
 - Limited memory BFGS [Liu & Nocedal '89]
- Majorization and bounding methods
 - Expectation-Maximization [Baum 1970, Dempster '77]
 - Generalized iterative scaling [Darroch & Ratcliff '72]
 - Majorization or majorize/minimize [deLeeuw & Heiser '77]
 - Quadratic lower bound principle [Bohning & Lindsay '88]
 - Improved iterative scaling [Berger et al. '97]
 - Extended Baum-Welch [Kanevsky et al. '08]



If cost function $\theta^* = \arg \min_{\theta} C(\theta)$ has no closed form solution Majorization uses with a surrogate Q with closed form update to monotonically minimize the cost from an initial θ_0

- Find bound $Q(\theta, \theta_i) \ge C(\theta)$ where $Q(\theta_i, \theta_i) = C(\theta_i)$
- Update $\theta_{i+1} = \arg \min_{\theta} Q(\theta, \theta_i)$

Repeat until converged





If cost function $\theta^* = \arg \min_{\theta} C(\theta)$ has no closed form solution Majorization uses with a surrogate Q with closed form update to monotonically minimize the cost from an initial θ_0

- Find bound $Q(\theta, \theta_i) \ge C(\theta)$ where $Q(\theta_i, \theta_i) = C(\theta_i)$
- Update $\theta_{i+1} = \arg \min_{\theta} Q(\theta, \theta_i)$

Repeat until converged





If cost function $\theta^* = \arg \min_{\theta} C(\theta)$ has no closed form solution Majorization uses with a surrogate Q with closed form update to monotonically minimize the cost from an initial θ_0

• Find bound $Q(\theta, \theta_i) \ge C(\theta)$ where $Q(\theta_i, \theta_i) = C(\theta_i)$

• Update
$$\theta_{i+1} = \arg \min_{\theta} Q(\theta, \theta_i)$$

Repeat until converged





- Central quantity to optimize in
 - Maximum likelihood and e-family [Pitman & Wishart '36]
 - Maximum entropy [Jaynes '57]
 - Conditional random fields [Lafferty, et al. '01]
 - Log-linear models [Darroch & Ratcliff '72]
 - Graphical models, HMMs [Jordan, et al. '99]
- Majorization preferred until [Wallach '03, Andrew & Gao '07]

Method	Iterations	LL Evaluations	Time (s)
IIS	≥ 150	≥ 150	≥ 188.65
Conjugate gradient (FR)	19	99	124.67
Conjugate gradient (PRP)	27	140	176.55
L-BFGS	22	22	29.72

The problem: loose & complicated bounds. Let's fix this!

・ロト ・ 日本・ 小田 ・ 小田 ・ 今日・



Consider log-linear model over discrete $y \in \Omega$ where $|\Omega| = n$

$$p(y|\theta) = \frac{1}{Z(\theta)}h(y)\exp\left(\theta^{\top}\mathbf{f}(y)\right)$$

- Parameters are vector $oldsymbol{ heta} \in \mathbb{R}^d$
- Features are $\mathbf{f}: \Omega \mapsto \mathbb{R}^d$ mapping each y to some vector
- Prior is $h: \Omega \mapsto \mathbb{R}^+$ a fixed non-negative measure
- Partition function ensures that $p(y|\theta)$ normalizes

$$Z(\theta) = \sum_{y} h(y) \exp(\theta^{\top} \mathbf{f}(y))$$

Problem: it's ugly to minimize, we much prefer quadratics

Partition Function Bound

The bound $\ln Z(\theta) \leq \ln z + \frac{1}{2}(\theta - \tilde{\theta})^{\top} \Sigma(\theta - \tilde{\theta}) + (\theta - \tilde{\theta})^{\top} \mu$ is tight at $\tilde{\theta}$ and holds for parameters given by



◆□▶ ◆□▶ ◆三▶ ◆三▶ 三三 少へ⊙

Bound Proof (Sketch)

Proof.

1) Start with bound $\log(e^{ heta}+e^{- heta})\leq c heta^2$ [Jaakkola & Jordan '99]

▲日▼▲□▼▲□▼▲□▼ □ ののの

- 2) Prove scalar bound via Fenchel dual using $\theta = \sqrt{\vartheta}$
- 3) Make bound multivariate $\log(e^{\theta^{\top}1} + e^{-\theta^{\top}1})$
- 4) Handle scaling of exponentials $\log(h_1 e^{\theta^{\top} \mathbf{f}_1} + h_2 e^{-\theta^{\top} \mathbf{f}_2})$
- 5) Add one term $\log(h_1 e^{\theta^\top \mathbf{f}_1} + h_2 e^{-\theta^\top \mathbf{f}_2} + h_3 e^{-\theta^\top \mathbf{f}_3})$
- 6) Repeat extension for *n* terms

The Bound as a Variation of Newton

$$\begin{array}{l} \begin{array}{l} \operatorname{Input} \tilde{\boldsymbol{\theta}}, \mathbf{f}(y), h(y) \, \forall y \in \Omega \\ \\ \operatorname{Init} z \to 0^+, \boldsymbol{\mu} = \mathbf{0}, \boldsymbol{\Sigma} = z \mathbf{I} \\ \end{array} \\ \operatorname{For each} y \in \Omega \left\{ \\ \alpha &= h(y) \exp(\tilde{\boldsymbol{\theta}}^\top \mathbf{f}(y)) \\ \mathbf{I} &= \mathbf{f}(y) - \boldsymbol{\mu} \\ \boldsymbol{\Sigma} + = \frac{\tanh(\frac{1}{2}\ln(\alpha/z))}{2\ln(\alpha/z)} \mathbf{I} \mathbf{I}^\top \\ \boldsymbol{\mu} + = \frac{\alpha}{z + \alpha} \mathbf{I} \\ z &+ = \alpha \end{array} \right\} \\ \end{array}$$

Computing the bound (left) and Newton's approximation (right) Both take $\mathcal{O}(nd^2)$ and update via $\theta \leftarrow \tilde{\theta} - \Sigma^{-1}\mu$ in $\mathcal{O}(d^3)$

・ロト ・ 理 ト ・ ヨ ト ・ ヨ ・



Maximum entropy (or generally) minimum relative entropy $\mathcal{RE}(p||h) = \sum_{y} p(y) \ln \frac{p(y)}{h(y)}$ subject to linear constraints

$$\min_{p} \mathcal{RE}(p \| h) \text{ s.t.} \sum_{y} p(y) \mathbf{f}(y) = \mathbf{0}, \sum_{y} p(y) \mathbf{g}(y) \ge \mathbf{0}$$

Its dual is the negative log-partition function (to be maximized):

$$-\ln Z(\boldsymbol{ heta}, \boldsymbol{artheta}) = -\ln \sum_{y} h(y) \exp\left(\boldsymbol{ heta}^{ op} \mathbf{f}(y) + \boldsymbol{artheta}^{ op} \mathbf{g}(y)
ight)$$

Maximum entropy is a natural application of the bound!

Conditional Random Fields (CRFs)

- Conditional random fields generalize maximum entropy
- Trained on *iid* data $\{(x_1, y_1), \ldots, (x_t, y_t)\}$
- Each CRF is a log-linear model

$$p(y|x_j, \theta) = \frac{1}{Z_{x_j}(\theta)} h_{x_j}(y) \exp(\theta^{\top} \mathbf{f}_{x_j}(y))$$

Regularized maximum likelihood objective function is

$$J(\boldsymbol{\theta}) = \sum_{j=1}^{t} \ln \frac{h_{x_j}(y_j)}{Z_{x_j}(\boldsymbol{\theta})} + \boldsymbol{\theta}^{\top} \mathbf{f}_{x_j}(y_j) - \frac{t\lambda}{2} \|\boldsymbol{\theta}\|^2$$
(1)

- Can even constrain the allowable heta inside convex Λ
- Permits ℓ_1 regularized CRFs, and other variants

Maximum Likelihood Algorithm for CRFs

Input
$$x_j, y_j$$
 and functions $h_{x_j}, \mathbf{f}_{x_j}$ for $j = 1, ..., t$
Input regularizer $\lambda \in \mathbb{R}^+$ and convex hull $\mathbf{\Lambda} \subseteq \mathbb{R}^d$
Initialize θ_0 anywhere inside $\mathbf{\Lambda}$ and set $\tilde{\theta} = \theta_0$
While not converged
For $j = 1$ to t compute bound for μ_j, Σ_j from $h_{x_j}, \mathbf{f}_{x_j}, \tilde{\theta}$
Set $\tilde{\theta} = \arg \min_{\theta \in \mathbf{\Lambda}} \sum_j \frac{1}{2} (\theta - \tilde{\theta})^\top (\Sigma_j + \lambda \mathbf{I}) (\theta - \tilde{\theta})$
 $+ \sum_j \theta^\top (\mu_j - \mathbf{f}_{x_j}(y_j) + \lambda \tilde{\theta})$
Output $\hat{\theta} = \tilde{\theta}$

Theorem

$$\begin{split} & If \, \|\mathbf{f}_{x_j}(y)\| \leq r \, get \, J(\hat{\theta}) - J(\theta_0) \geq (1-\epsilon) \max_{\theta \in \mathbf{\Lambda}} (J(\theta) - J(\theta_0)) \\ & within \, \left\lceil \ln\left(1/\epsilon\right) / \ln\left(1 + \frac{\lambda \log n}{2r^2 n}\right) \right\rceil \, steps \end{split}$$

Convergence Proof

Proof.



Figure: Quadratic bounding sandwich. Compare upper and lower bound curvatures to bound maximum # of iterations.

Mixture Models Mixture Models Experiments

Bounding Graphical Models with Large n



- Each iteration is $\mathcal{O}(tnd^2)$, but what if *n* is large?
- Graphical model: a bipartite factor graph G representing a distribution p(Y) where $Y = \{y_1, \ldots, y_n\}$ and $y_i \in \mathbb{Z}$
- p(Y) factorizes as product of $\{\psi_1, \ldots, \psi_C\}$ functions (squares) over $\{Y_1, \ldots, Y_C\}$ subsets of variables (circles)

$$p(y_1,\ldots,y_n) = \frac{1}{Z}\prod_{c\in C}\psi_c(Y_c)$$

• E.g. $p(y_1, \ldots, y_6) = \psi(y_1, y_2)\psi(y_2, y_3)\psi(y_3, y_4, y_5)\psi(y_4, y_5, y_6)$

Bounding Graphical Models with Large n



- Instead of enumerating over all *n*, exploit graphical model
- Build junction tree and run a *Collect* algorithm
- Already used for computing $Z(\theta)$ and $Z'(\theta)$ efficiently
- Bound needs $\mathcal{O}(td^2\sum_c |Y_c|)$ rather than $\mathcal{O}(td^2n)$
- For an HMM, this is $\mathcal{O}(TM^2)$ instead of $\mathcal{O}(M^T)$

Optimization Partition Bound Graphical Models Low Rank Bound Mixture Models Mixture Models Experiments

Bounding Graphical Models with Large n

for
$$c = 1, ..., m$$
 {
 $Y_{both} = Y_c \cap Y_{pa(c)}; Y_{solo} = Y_c \setminus Y_{pa(c)}$
for each $u \in Y_{both}$ {
initialize $z_{c|x} \leftarrow 0^+$, $\mu_{c|x} = \mathbf{0}, \Sigma_{c|x} = z_{c|x}\mathbf{I}$
for each $v \in Y_{solo}$ {
 $w = u \otimes v; \quad \alpha_w = h_c(w)e^{\tilde{\theta}^\top \mathbf{f}_c(w)} \prod_{b \in ch(c)} z_{b|w}$
 $\mathbf{I}_w = \mathbf{f}_c(w) - \mu_{c|u} + \sum_{b \in ch(c)} \mu_{b|w}$
 $\Sigma_{c|u} + = \sum_{b \in ch(c)} \Sigma_{b|w} + \frac{\tanh(\frac{1}{2}\ln(\frac{\alpha_w}{z_{c|u}}))}{2\ln(\frac{\alpha_w}{z_{c|u}})} \mathbf{I}_w \mathbf{I}_w^\top$
 $\mu_{c|u} + = \frac{\alpha_w}{z_{c|u} + \alpha_w} \mathbf{I}_w; \quad z_{c|u} + = \alpha_w$ }}

Low Rank Bound for Large d

- Naive bounding takes $\mathcal{O}(tnd^2)$, inverting takes $\mathcal{O}(d^3)$
- To match gradient methods and LBFGS, need $\mathcal{O}(tnd)$
- Consider a rank 1 update: $\Sigma += \frac{\tanh(\frac{1}{2}\ln(\alpha/z))}{2\ln(\alpha/z)} \mathbf{II}^{\top}$
- As in LBFGS, use rank-k storage $\boldsymbol{\Sigma} = \boldsymbol{\mathsf{VSV}}^{ op} + \boldsymbol{\mathsf{D}}$
- ullet Each rank 1 update on Σ is projected on $oldsymbol{V}$
- Top k eigenvectors are kept with updated eigenvalues in S
- Remaining residual is absorbed into diagonal **D**
- By Jensen inequality on diagonal **D**, low-rank is still a bound

• Avoid $\mathcal{O}(d^3)$ inversion in $\theta = \tilde{\theta} - \Sigma^{-1}\mu$: use Woodbury formula, $\Sigma^{-1} = \mathbf{D}^{-1} + \mathbf{D}^{-1}\mathbf{V}^{\top}(\mathbf{S}^{-1} + \mathbf{V}\mathbf{D}^{-1}\mathbf{V}^{\top})^{-1}\mathbf{V}\mathbf{D}^{-1}$ with only $\mathcal{O}(k^3)$ work

◆□▶ ◆□▶ ◆ □▶ ★ □▶ = □ ● の < @

Low Rank Bound for Large d in O(tndk)

for each
$$t \{ \text{set } z \to 0^+; \mu = \mathbf{0};$$

for each $y \{$

$$\alpha = h_t(y)e^{\tilde{\theta}^\top \mathbf{f}_t(y)}; \mathbf{r} = \frac{\sqrt{\tanh(\frac{1}{2}\log(\frac{\alpha}{z}))}}{\sqrt{2\log(\frac{\alpha}{z})}} (\mathbf{f}_t(y) - \mu);$$
For $i = 1, \dots, k : \mathbf{p}(i) = \mathbf{r}^\top \mathbf{V}(i, \cdot); \mathbf{r} = \mathbf{r} - \mathbf{p}(i)\mathbf{V}(i, \cdot);$
For $i = 1, \dots, k : \text{For } j = 1, \dots, k : \mathbf{S}(i, j) = \mathbf{S}(i, j) + \mathbf{p}(i)\mathbf{p}(j);$

$$\mathbf{Q}^\top \mathbf{A} \mathbf{Q} = \operatorname{svd}(\mathbf{S}); \mathbf{S} \leftarrow \mathbf{A}; \mathbf{V} \leftarrow \mathbf{Q} \mathbf{V};$$

$$\mathbf{s} = [\mathbf{S}(1, 1), \dots, \mathbf{S}(k, k), \|\mathbf{r}\|^2]^\top; \quad \tilde{k} = \arg\min_{i=1,\dots,k+1} \mathbf{s}(i);$$
if $(\tilde{k} \le k) \{ \mathbf{D} = \mathbf{D} + \mathbf{S}(\tilde{k}, \tilde{k})\mathbf{1}^\top |\mathbf{V}(j, \cdot)| \operatorname{diag}(|\mathbf{V}(k, \cdot)|);$

$$\mathbf{S}(\tilde{k}, \tilde{k}) = \|\mathbf{r}\|^2; \quad \mathbf{r} = \|\mathbf{r}\|^{-1}\mathbf{r}; \quad \mathbf{V}(k, \cdot) = \mathbf{r}; \}$$
else
$$\{ \mathbf{D} = \mathbf{D} + \mathbf{1}^\top |\mathbf{r}| \operatorname{diag}(|\mathbf{r}|); \} \}$$

$$\mu + = \frac{\alpha}{z + \alpha} (\mathbf{f}_t(y) - \mu); \quad z + = \alpha; \}$$

Mixture Models and Latent Likelihood

- Bounding also simplifies mixture models with hidden variables
- Allows mixtures of Gaussians, HMMs, latent graphical models
- Assume data-set is generated by a conditional distribution

$$p(y|x,\Theta) = \frac{\sum_{m} p(x, y, m|\Theta)}{\sum_{y,m} p(x, y, m|\Theta)}$$

It is natural to maximize incomplete likelihood

$$L(\Theta) = \prod_{j=1}^{t} p(y_j | x_j, \Theta) = \prod_{j=1}^{t} \frac{\sum_m p(x_j, y_j, m | \Theta)}{\sum_{y, m} p(x_j, y, m | \Theta)}$$

• Assume exponential family mixture components (Gaussian, multinomial, Poisson, Laplace)

$$p(x|y, m, \Theta) = h(x) \exp\left(\theta_{y,m}^{\top} \phi_{y,m}(x) - a_{y,m}(\theta_{y,m})\right)$$

Mixture Models and Latent Likelihood

- Latent CRFs are just log-linear mixtures [Quattoni '07]
- When a CRF has hidden variable m, the latent likelihood is

$$L(\boldsymbol{\theta}) = \prod_{j=1}^{t} \frac{\sum_{m} \exp\left(\boldsymbol{\theta}^{\top} \mathbf{f}_{j,y_{j},m}\right)}{\sum_{y,m} \exp\left(\boldsymbol{\theta}^{\top} \mathbf{f}_{j,y,m}\right)}.$$

- Vectors **f** are concatenations of $\phi_{y,m}(x)$ sufficient statistics
- Apply Jensen to numerator and our bound to denominator
- Get an auxiliary function $L(heta) \geq Q(heta, ilde{ heta})$
- Maximizing $Q(heta, ilde{ heta})$ is just a matrix inverse
- Like CEM or conditional variant of EM [J & Pentland '00]
- If |m| = 1, reduces back to usual conditional random field

Mixture Models and Latent Likelihood



Two Gaussian mixture model trained for maximum conditional likelihood (left) and maximum likelihood (right)

▲日▼▲□▼▲□▼▲□▼ □ ののの

Experiments

Experiments - Classification and Structured Prediction

Data-set	SRB	CT Tumors		s	Text		SecStr		CoNLL		PennTree	
Size	<i>n</i> = 4		n = 26 n =		2	<i>n</i> = 2		<i>m</i> = 9		<i>m</i> = 45		
	t = 83		<i>t</i> = 30	8	t = 1500		t = 83679		t = 1000		t = 1000	
	d = 9	9236	d = 3902	260	d = 23	922	d = 63	32	<i>d</i> = 336	15	d = 141	75
	$\lambda =$	10^{1}	$\lambda = 10$	$\lambda = 10^1$ $\lambda = 10^2$		$\lambda = 10^1$		$\lambda = 10^1$		$\lambda = 10^1$		
Algorithm	time	iter	time	iter	time	iter	time	iter	time	iter	time	iter
LBFGS	6.10	42	3246.83	8	15.54	7	881.31	47	25661.54	17	62848.08	7
SD	7.27	43	18749.15	53	153.10	69	1490.51	79	93821.72	12	156319.31	12
CG	40.61	100	14840.66	42	57.30	23	667.67	36	88973.93	23	76332.39	18
Bound	3.67	8	1639.93	4	6.18	3	27.97	9	16445.93	4	27073.42	2

Table: Time in seconds and iterations to match LBEGS solution for logistic regression (on SRBCT, Tumors, Text and SecStr data-sets where *n* is the number of classes) and Markov CRFs (on CoNLL and PennTree data-sets, where m is the number of classes). Here, t is the total number of samples (training and testing), d is the dimensionality of the feature vector and λ is the cross-validated regularization setting.

Experiments - Testing Latent Likelihood

Data-set	ion	bupa	hepatitis	wine	SRBCT
Algorithm	m = 3	m=2	m = 2	m = 3	m = 4
BFGS	-5.88	-21.78	-5.28	-1.79	-6.06
SD	-5.56	-21.74	-5.14	-1.37	-5.61
CG	-5.57	-21.81	-4.84	-0.95	-5.76
Newton	-5.95	-21.85	-5.50	-0.71	-5.54
Bound	-4.18	-19.95	-4.40	-0.48	-0.11

Table: Test log-likelihood at convergence for ion, bupa, hepatitis, wine and SRBCT data-sets.



Figure: Convergence of latent likelihood over time for several data-sets. (a) < ≥ > < ≥ > - 34



- Majorization was non-competitive due to slow & loose bounds
- We derived *simple* quadratic bound on the partition function
- Makes majorization competitive with state-of-the-art
- Bound is efficient for graphical models and large n
- Low-rank bound is efficient for large dimensionality d
- Yields fast and monotonically convergent majorization
- Used for maximum entropy, CRFs and latent likelihood
- Current work: HMMs, stochastic bounds, loopy graphs, deep belief networks, distributed optimization