

Dimension reduction

COMS 4771 Fall 2023

Linear dimension reduction

Dimension reduction: map feature vectors from \mathbb{R}^d to \mathbb{R}^k with $k < d$

- ▶ Reduce storage requirements for dataset
- ▶ Improve understandability of individual data points
- ▶ Improve performance of learning algorithms on dataset
- ▶ ...

Many methods are linear: i.e., based on linear map $\varphi: \mathbb{R}^d \rightarrow \mathbb{R}^k$

This lecture: unsupervised methods for dimension reduction

Throughout this lecture, $X = (X_1, \dots, X_d)$ is a random vector

e.g., $X =$ data point drawn uniformly at random from \mathcal{S}

Axis-aligned embeddings

Axis-aligned embeddings:

- ▶ Let $\varphi(x) \in \mathbb{R}^k$ keep a subset of k features x_i , throw away the rest

Question: Which features to keep?

- ▶ Simple heuristic: Choose the k most “informative” features

Sort features by variance

$$\text{var}(X_{(1)}) \geq \dots \geq \text{var}(X_{(d)})$$

and choose $\varphi(x) = (x_{(1)}, \dots, x_{(k)})$

Suppose only k features have non-negligible variance

$$\text{var}(X_{(1)}) \geq \dots \geq \text{var}(X_{(k)}) \gg \text{var}(X_{(k+1)}) \approx \dots \approx \text{var}(X_{(d)}) \approx 0$$

And $\varphi(x) = (x_{(1)}, \dots, x_{(k)}) \in \mathbb{R}^k$

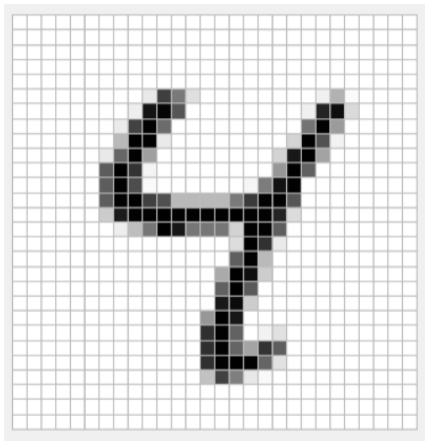
For affine function $w^\top x + b$, we have

$$w^\top X + b \approx$$

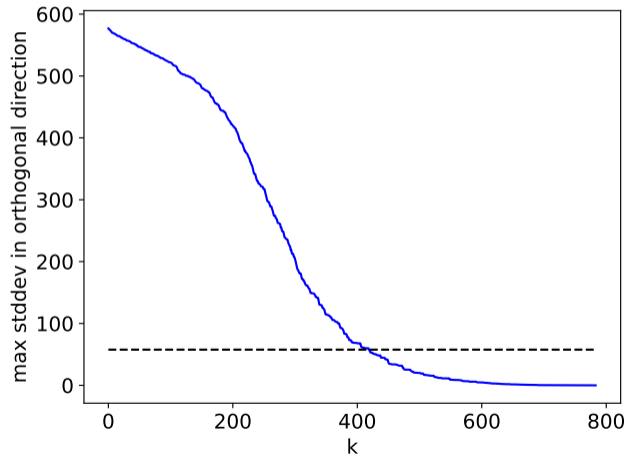
Therefore, this is close to $\tilde{w}^\top \varphi(X) + \tilde{b}$ for some $\tilde{w} \in \mathbb{R}^k$ and $\tilde{b} \in \mathbb{R}$

Example: MNIST dataset of handwritten digit images

- ▶ 784 features corresponding to pixel intensity values (from $\{0, 1, \dots, 255\}$)



Vertical axis: $\max_{\beta \in \mathbb{R}^{d-k}} \frac{\text{stddev}(\beta_{k+1}X_{(k+1)} + \dots + \beta_d X_{(d)})}{\|\beta\|}$



Can we do better than “axis-aligned embeddings”?

- ▶ Maybe there is a better way to choose which variables to keep?
- ▶ Retained features could contain a lot of redundancy!
- ▶ Can possibly reduce dimension even further by accounting for covariance between features

Covariance matrices

Covariance matrix $\text{cov}(X)$ of a random vector $X = (X_1, \dots, X_d)$:

- ▶ $d \times d$ matrix whose (i, j) -th entry is $\text{cov}(X_i, X_j)$
- ▶ Matrix notation:

$$\text{cov}(X) = \mathbb{E}[(X - \mathbb{E}(X))(X - \mathbb{E}(X))^{\top}]$$

- ▶ $\text{cov}(X)$ “encodes” covariance between all linear functions of X

Consider linear function $f(x) = \alpha^T x$, given by some $\alpha \in \mathbb{R}^d$

- ▶ If α is a unit vector (i.e., $\|\alpha\| = 1$), then $\alpha^T x$ is the “coordinate” of the orthogonal projection of x to the line spanned by α
- ▶ The “coordinate” $\alpha^T x$ is often referred to as the “projection of x in direction α ”, even though this is not technically correct

► What is the mean of $\alpha^T X$?

► What is the variance of $\alpha^T X$?

- ▶ What is the covariance between $\alpha^T X$ and $\beta^T X$?

Example: Dartmouth student data

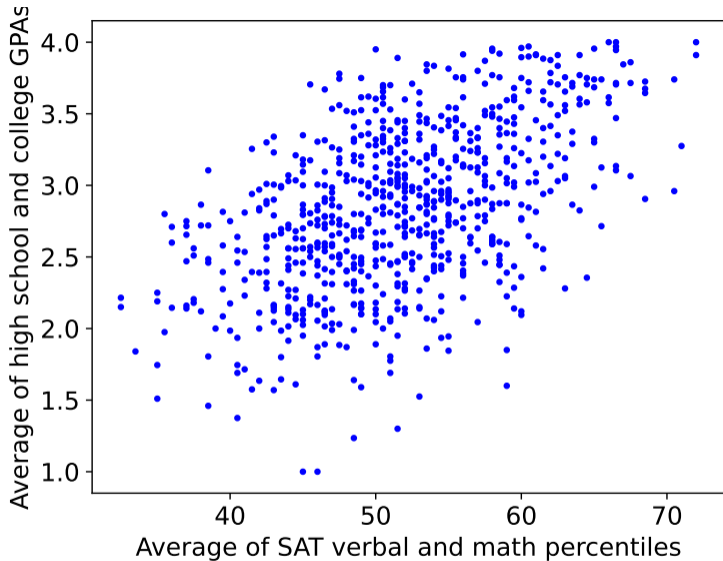
- ▶ $x_1 =$ SAT verbal percentile, $x_2 =$ SAT math percentile,
 $x_3 =$ high school GPA, $x_4 =$ (first year) college GPA
- ▶ $X =$ data point drawn uniformly at random from dataset

$$\text{cov}(X) = \begin{bmatrix} 69.8 & 33.8 & 1.74 & 2.71 \\ 33.8 & 72.3 & 1.76 & 2.43 \\ 1.74 & 1.76 & 0.29 & 0.22 \\ 2.71 & 2.43 & 0.22 & 0.56 \end{bmatrix}$$

- ▶ Define random variables Y and Z :

$$Y = \frac{1}{2}(\text{SAT verbal} + \text{SAT math})$$

$$Z = \frac{1}{2}(\text{high school GPA} + \text{college GPA})$$



Using $\text{cov}(X)$, can compute $\text{cor}(Y, Z)$:

$$\text{var}(Y) = \alpha^T \text{cov}(X) \alpha = 52.4$$

$$\text{var}(Z) = \beta^T \text{cov}(X) \beta = 0.32$$

$$\text{cov}(Y, Z) = \alpha^T \text{cov}(X) \beta = 2.16$$

$$\text{cor}(Y, Z) = \frac{\text{cov}(Y, Z)}{\sqrt{\text{var}(Y) \text{var}(Z)}} = 0.52$$

where

$$\alpha = \underline{\hspace{2cm}}$$

$$\beta = \underline{\hspace{2cm}}$$

Review of eigenvalues and eigenvectors

- ▶ Every symmetric $d \times d$ matrix M has d real eigenvalues, conventionally numbered in non-increasing order

$$\lambda_1 \geq \dots \geq \lambda_d$$

- ▶ Because M is symmetric, it is always possible to find d corresponding eigenvectors that form an orthonormal basis for \mathbb{R}^d :

$$v_1, \dots, v_d \in \mathbb{R}^d$$

such that

$$Mv_i = \lambda_i v_i$$

and

$$v_i^\top v_j = \begin{cases} 1 & \text{if } i = j \\ 0 & \text{otherwise} \end{cases}$$

Eigendecomposition of M

$$M = \sum_{i=1}^d \lambda_i v_i v_i^T$$

For rest of lecture, let $\text{cov}(X)$ have eigendecomposition

$$\text{cov}(X) = \sum_{i=1}^d \lambda_i v_i v_i^\top$$

with $\lambda_1 \geq \dots \geq \lambda_d$ and v_1, \dots, v_d orthonormal

Variance maximizing direction

“Variance of X in direction α ”:

$$\text{var}\left(\frac{1}{\|\alpha\|}\alpha^\top X\right) = \frac{\alpha^\top \text{cov}(X)\alpha}{\|\alpha\|^2}$$

Question: In which direction α does X have the highest variance?

$$\max_{\alpha \in \mathbb{R}^d \setminus \{0\}} \frac{\alpha^\top \text{cov}(X)\alpha}{\|\alpha\|^2}$$

Answer: $\alpha = v_1$ —i.e., eigenvector of $\text{cov}(X)$ corresponding to largest eigenvalue (a.k.a. top eigenvector)

Upshot: If you want to reduce to dimension $k = 1$, use direction of the top eigenvector of $\text{cov}(X)$

Example: MNIST (just the 8's); 10 images sorted by "coordinate" along v_1



Principal components analysis

What we want: minimize variance of X in directions that are “thrown away”

For $k = 1$, goal is captured by following problem:

$$\min_{\alpha \in \mathbb{R}^d} \max_{\substack{\beta \in \mathbb{R}^d \setminus \{0\}, \\ \beta \perp \alpha}} \frac{\beta^\top \text{cov}(X) \beta}{\|\beta\|^2}$$

Solution also is given by $\alpha = v_1$

This fact is a special case of the “Courant min-max principle”

- ▶ For $\alpha = v_1$,

$$\max_{\substack{\beta \in \mathbb{R}^d \setminus \{0\}, \\ \beta \perp \alpha}} \frac{\beta^\top \text{cov}(X)\beta}{\|\beta\|^2} = \underline{\hspace{2cm}}$$

- ▶ For any other α :

Courant min-max principle says

$$\min_{\substack{\mathcal{W} \subseteq \mathbb{R}^d, \\ \dim(\mathcal{W})=k}} \max_{\substack{\beta \in \mathbb{R}^d \setminus \{0\}, \\ \beta \perp \mathcal{W}}} \frac{\beta^\top \text{cov}(X)\beta}{\|\beta\|^2} = \text{-----}$$

and this is achieved by the subspace $\mathcal{W} = \text{span}\{v_1, \dots, v_k\}$ spanned by top- k eigenvectors of $\text{cov}(X)$

Principal components analysis (PCA): dimension reduction method that, for target dimension k , uses the linear map

$$\varphi(x) = (v_1^\top x, \dots, v_k^\top x)$$

based on the top- k eigenvectors of $\text{cov}(X)$

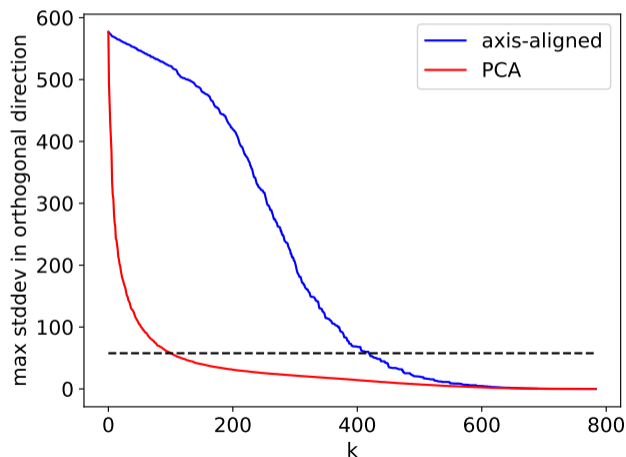
- ▶ $\varphi(x)$ gives the “coordinates” of the orthogonal projection of x to span of v_1, \dots, v_k , a.k.a. the dimension- k PCA projection
- ▶ Also

$$\text{cov}(\varphi(X)_i, \varphi(X)_j) =$$

So new “variables” in $\varphi(X)$ are uncorrelated

MNIST: What subspace dimension k is needed so worst standard deviation in an orthogonal direction is at most $0.1 \times \lambda_1$?

- ▶ Axis-aligned embeddings: $k = 419$; PCA embeddings: $k = 101$



Given $\varphi(x) \in \mathbb{R}^k$ (from PCA), along with v_1, \dots, v_k , can obtain d -dimensional “reconstruction” of x :

$$\sum_{i=1}^k \varphi(x)_i v_i$$

(orthogonal projection of x to the subspace spanned by v_1, \dots, v_k)

MNIST

original

$k = 25$

$k = 50$

$k = 75$

$k = 100$



Matrix approximation

PCA (on finite dataset) is related to singular value decomposition of $n \times d$ matrix

$$A = \begin{bmatrix} \leftarrow & (x^{(1)})^\top & \rightarrow \\ & \vdots & \\ \leftarrow & (x^{(n)})^\top & \rightarrow \end{bmatrix}$$

Every matrix A has a singular value decomposition (SVD): decomposition of A into the sum of r rank-1 matrices

$$A = \sum_{i=1}^r s_i u^{(i)} (v^{(i)})^\top$$

where

- ▶ $r = \text{rank}(A)$
- ▶ $s_1 \geq \dots \geq s_r > 0$ as positive real numbers (singular values of A)
- ▶ $u^{(1)}, \dots, u^{(r)}$ is ONB for $\text{CS}(A)$ (left singular vectors of A)
- ▶ $v^{(1)}, \dots, v^{(r)}$ is ONB for $\text{CS}(A^\top)$ (right singular vectors of A)

Matrix form of SVD:

$$A = \underbrace{\begin{bmatrix} \uparrow & & \uparrow \\ u^{(1)} & \dots & u^{(r)} \\ \downarrow & & \downarrow \end{bmatrix}}_U \underbrace{\begin{bmatrix} s_1 & & \\ & \ddots & \\ & & s_r \end{bmatrix}}_S \underbrace{\begin{bmatrix} \leftarrow & (v^{(1)})^\top & \rightarrow \\ & \vdots & \\ \leftarrow & (v^{(r)})^\top & \rightarrow \end{bmatrix}}_{V^\top}$$

Computation: `numpy.linalg.svd`

Rank- k (truncated) SVD: keep only the first $k \leq r$ components of the SVD

$$A^{(k)} = \sum_{i=1}^k s_i u^{(i)} (v^{(i)})^\top$$

In matrix form:

$$A^{(k)} = \underbrace{\begin{bmatrix} \uparrow & & \uparrow \\ u^{(1)} & \dots & u^{(k)} \\ \downarrow & & \downarrow \end{bmatrix}}_{U^{(k)}} \underbrace{\begin{bmatrix} s_1 & & \\ & \ddots & \\ & & s_k \end{bmatrix}}_{S^{(k)}} \underbrace{\begin{bmatrix} \leftarrow & (v^{(1)})^\top & \rightarrow \\ & \vdots & \\ \leftarrow & (v^{(k)})^\top & \rightarrow \end{bmatrix}}_{(V^{(k)})^\top}$$

Eckart-Young Theorem: If $k \leq \text{rank}(A)$, then $A^{(k)} = \sum_{i=1}^k s_i u^{(i)} (v^{(i)})^\top$ from rank- k SVD has smallest sum-of-squared errors

$$\sum_{i=1}^n \sum_{j=1}^d (A_{i,j} - \tilde{A}_{i,j})^2$$

among all $n \times d$ matrices \tilde{A} of rank k

Connection to PCA: Let X be random vector with uniform distribution over $\{x^{(1)}, \dots, x^{(n)}\}$ (and assume A is row-centered, so $\frac{1}{n} \sum_{i=1}^n x^{(i)} = 0$)

▶ Then $\text{cov}(X) =$ _____

▶ Moreover,

$$A^T A = \text{_____}$$

▶ Non-zero eigenvalues of $\text{cov}(X)$ are _____

▶ Corresponding eigenvectors of $\text{cov}(X)$ are _____

Statistical model: A is $n \times d$ matrix of independent random variables, with

$$A_{i,j} \sim N(H_{i,j}, \sigma^2)$$

where H is $n \times d$ matrix with $\text{rank} \leq k$ (the “parameter” of this model)

Maximum likelihood estimator of H : _____

